

# Threshold Profiling for Wideband Ranging

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**Abstract**—This letter establishes a methodology to design threshold profiles for wideband ranging systems. Differently from conventional methods using a single threshold value that is designed based on the signal detection and false-alarm probability requirements, we propose a threshold profile that is designed based also on the ranging error requirement. The proposed method relies only on channel statistics without requiring channel estimation. A case study shows that, compared to conventional methods, the proposed method for threshold profiling significantly improves the performance in terms of false-alarm probability, detection probability, and ranging error.

**Index Terms**—Energy detection, threshold profiling, time-of-arrival estimation, wideband ranging.

## I. INTRODUCTION

NETWORK LOCALIZATION [1] is a key enabler for a variety of emerging applications including autonomous vehicles, smart cities, logistics, public safety, environmental monitoring, and social networks [2]–[10]. Range-based network localization relies on measurements of distances between wireless nodes [11]–[13]. In particular, ranging techniques based on time-of-arrival (TOA) estimation are widely adopted, especially for high-accuracy wideband localization in harsh propagation environments (e.g., indoor). Conventional techniques for TOA estimation rely on threshold-based detection, in which the choice of the threshold is crucial for detection probability, false-alarm probability, and ranging error.

The design of threshold-based detectors is challenging due to impairments in wireless channels such as multipath, obstructions, and noise. If the knowledge of the instantaneous channel (multipath amplitudes and delays) is available, then the matched filter enables the optimal coherent detection [11]. However, non-coherent detectors, which require only the knowledge of channel

statistics (power decay and noise floor), are preferred to coherent detectors for their lower complexity.

A popular noncoherent detector is the energy detector (ED), which determines the TOA by comparing energy samples collected in consecutive dwell intervals with a threshold [14]–[17]. In conventional approaches, the threshold is the same for all dwell intervals and is chosen to provide a constant false-alarm rate (CFAR) [18]. Specifically, the threshold is determined by accounting only for the noise without considering the wireless environment [19]–[22]. Accounting for the wireless environment would result in a threshold profile (i.e., a different threshold for each dwell interval) instead of a threshold value. For example, a delay-dependent threshold was obtained via simulation in [23].

This letter proposes a method for designing threshold profiles for wideband ranging systems. The proposed threshold profile meets the requirements on ranging error in addition to those on false-alarm probability and detection probability. In particular, a tractable expression for the threshold profile is derived for wideband signals. The performance improvement of threshold profiling with respect to fixed thresholds is quantified in a case of study.

## II. RANGING SYSTEM MODEL

This section presents the ranging system model: first, the statistical distribution of the energy samples is provided; and then, the range inference method is described.

### A. Energy Samples

Consider a transmitter at position  $p_t$  and a receiver at position  $p_r$ . The transmitter emits a sequence containing  $N_{sr}$  replicas of a signal  $s(t)$ , with repetition frequency  $f_{sr} = 1/T_{sr}$ . At the receiver, the replicas are noncoherently accumulated. The aim of a ranging system is to detect the presence of the signal  $s(t)$  from the received waveform and to estimate its TOA  $\tau$  with respect to a reference time  $t_0$ .<sup>1</sup> The reference time  $t_0$  can be the time at which the signal was transmitted (e.g., localization or radar networks based on TOA) or be the time shared among several receivers (e.g., localization or radar networks based on time-difference-of-arrival).

Fig. 1 illustrates the signal processing elements for an ED-based ranging system after bandpass filtering for noise reduction

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<sup>1</sup>Range and TOA are used interchangeably throughout this letter since the former is a bijective function of the latter.

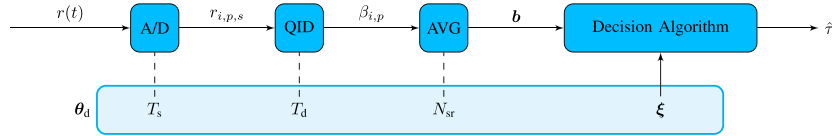


Fig. 1. Ranging system based on energy detection.

(and after clutter mitigation in radar networks).<sup>2</sup> In particular, the filtered signal  $r(t)$  is first sampled by an analog-to-digital converter (A/D) with sampling interval  $T_s$ . The waveform samples are then processed by a quadrature integrate and dump (QID) block that squares and integrates the samples over a dwell time  $T_d$  to obtain  $N_{\text{bin}} = \lfloor T_{\text{obs}}/T_d \rfloor$  energy bins within the observation time  $T_{\text{obs}}$ , where  $\lfloor x \rfloor$  is the greatest integer less than or equal to  $x$ . Each bin is obtained from  $N_{\text{sb}} = \lfloor T_d/T_s \rfloor$  received samples. The  $i$ th bin corresponds to the integration interval  $\mathcal{T}_i = (T_d i, T_d (i+1)]$  with  $i \in \mathcal{B} = \{0, 1, \dots, N_{\text{bin}} - 1\}$ . The energy bins are then processed by an averaging (AVG) block over the  $N_{\text{sr}}$  replicas to increase the signal-to-noise ratio (SNR). The output of the AVG block results in a vector of energy bins  $\mathbf{b} = [b_0, b_1, \dots, b_{N_{\text{bin}}-1}]$ . A ranging algorithm chooses a bin index  $\hat{i}$  based on the energy bin vector  $\mathbf{b}$ . The TOA estimate is then determined as a value inside the interval  $\mathcal{T}_{\hat{i}}$  according to a function  $\hat{\tau} = g(\hat{i})$  (e.g.,  $g(\hat{i}) = \hat{i}T_d + T_d/2$ ).

Each element  $b_i$  of  $\mathbf{b}$  is the instantiation of a random variable (RV)  $b_i$  with probability distribution depending on the transmitted signal, the true TOA  $\tau$ , the wireless channel, and the ED parameters. Let  $\boldsymbol{\theta}_b = [\tau, \boldsymbol{\theta}_h, \boldsymbol{\theta}_d]$  denotes the parameter vector, with  $\boldsymbol{\theta}_h$  representing the wireless channel parameters (including the amplitude and arrival time of multipath components) and  $\boldsymbol{\theta}_d$  representing the ED parameters (including dwell time and threshold values). The normalized bin values  $b_i N_{\text{sr}}/\sigma^2$ , with  $i \in \mathcal{B}$ , conditioned on  $\boldsymbol{\theta}_b$  are independent and distributed as noncentral chi-squared RVs with  $N_{\text{sr}}N_{\text{sb}}$  degrees of freedom [14], i.e.,

$$b_i \frac{N_{\text{sr}}}{\sigma^2} \Big|_{\boldsymbol{\theta}_b} \overset{\sim}{\sim} \chi_{N_{\text{sr}}N_{\text{sb}}}^2(\lambda_i). \quad (1)$$

The parameter  $\lambda_i$  is the noncentrality parameter that depends on  $\boldsymbol{\theta}_b$  and is given by

$$\lambda_i = \sum_{p=0}^{N_{\text{sr}}-1} \sum_{s=0}^{N_{\text{sb}}-1} \frac{u_{i,p,s}^2}{\sigma^2} \quad (2)$$

where  $u_{i,p,s} = u(t_{i,p,s})$  is the sample of the signal after propagating through a wireless channel with impulse response  $h(t)$ ;  $t_{i,p,s} = iT_d + pT_{\text{sr}} + sT_s$  is the sampling instant; and  $\sigma^2$  is the variance of the zero-mean Gaussian noise.

### B. Range Inference

Several TOA estimation algorithms can be found in the literature when the input is a vector of energy bins, including the threshold crossing search (TCS), jump back and search forward (JBSF), and serial backward search (SBS) [11]. Most of them require the comparison of each bin value with a threshold. In

the following, the TCS algorithm is used as a case study to illustrate the design of the threshold profile, a similar approach can be used for other threshold-based algorithms.

The TCS algorithm first searches for any bin value  $b_i$  that crosses a threshold  $\xi_i$  for all  $i \in \mathcal{B}$ . The algorithm then selects, if  $\mathcal{C}_{\text{th}} = \{\exists i \in \mathcal{B} \text{ s.t. } b_i > \xi_i\}$  occurs, the bin index  $\hat{i}$  as the smallest  $i$  for which  $b_i > \xi_i$ . When  $\mathcal{C}_{\text{th}}$  does not occur, the signal is undetected and the TOA is not estimated [25].

Consider the parameter vector  $\boldsymbol{\theta}$  such that the corresponding  $\boldsymbol{\lambda} \neq \mathbf{0}$ , where  $\boldsymbol{\lambda} = [\lambda_0, \lambda_1, \dots, \lambda_{N_{\text{bin}}-1}]$  is the vector of noncentrality parameters. The detection probability is given by

$$P_d(\boldsymbol{\theta}_b) = 1 - \prod_{i \in \mathcal{B}} F_{b_i}(\xi_i | \boldsymbol{\theta}_b) \quad (3)$$

where  $F_{b_i}(\cdot | \boldsymbol{\theta}_b)$  is the cumulative distribution function (CDF) of the RV  $b_i$  that depends on the parameter vector  $\boldsymbol{\theta}_b$ . The false-alarm event occurs when, in the absence of the transmitted signal, the signal is incorrectly detected.<sup>3</sup> Consider the parameters vector  $\boldsymbol{\theta}$  such that the corresponding  $\boldsymbol{\lambda} = \mathbf{0}$ . The false-alarm probability is given by

$$P_{\text{fa}}(\boldsymbol{\theta}_b) = 1 - \prod_{i \in \mathcal{B}} F_{b_i}(\xi_i | \boldsymbol{\theta}_b). \quad (4)$$

The ranging error can be characterized in several ways. For a true TOA  $\tau \in \mathcal{T}_i$ , the minimum ranging error is obtained when the bin  $i = \lfloor \tau/T_d \rfloor$  is selected, for which  $|\hat{\tau} - \tau| \leq T_d$  for any  $g(\cdot)$ . Considering the parameters vector  $\boldsymbol{\theta}$  for  $\tau \in \mathcal{T}_i$ , the probability of selecting the correct bin  $i$  is

$$P_{c,i}(\boldsymbol{\theta}_b) = \left[1 - F_{b_i}(\xi_i | \boldsymbol{\theta}_b)\right] \prod_{j \in \mathcal{I}_n(i)} F_{b_j}(\xi_j | \boldsymbol{\theta}_b) \quad (5)$$

where  $\mathcal{I}_n(m) \triangleq \mathcal{B} \cap \{m-n, m-n+1, \dots, m-1\}$ , with  $m > n$  and  $m > 0$ , and is empty for  $n \leq 0$ .

### III. DESIGN OF THRESHOLD PROFILE

The design of the threshold profile vector  $\boldsymbol{\xi} = [\xi_0, \xi_1, \dots, \xi_{N_{\text{bin}}-1}]$  is crucial for ranging error. Consider the  $N_{\text{bin}}$  hypothesis tests for a given channel instantiation  $\boldsymbol{\theta}$

$$\mathcal{H}_{0i} : b_i \frac{N_{\text{sr}}}{\sigma^2} \Big|_{\boldsymbol{\theta}_b} \overset{\sim}{\sim} \chi_{N_{\text{sr}}N_{\text{sb}}}^2(0) \quad (6a)$$

$$\mathcal{H}_{1i} : b_i \frac{N_{\text{sr}}}{\sigma^2} \Big|_{\boldsymbol{\theta}_b} \overset{\sim}{\sim} \chi_{N_{\text{sr}}N_{\text{sb}}}^2(\lambda_i \neq 0) \quad (6b)$$

where  $i \in \mathcal{B}$ ,  $\mathcal{H}_{0i}$ , and  $\mathcal{H}_{1i}$  are the hypotheses that the  $i$ th bin contains noise only or also the transmitted signal, respectively. The optimal test rejects  $\mathcal{H}_{0i}$  with test statistics  $b_i > \xi_i$ , and TCS selects the smallest bin index  $i$  for which this happens.

<sup>2</sup>Note that the results of this letter can also be applied to different versions of ED-based ranging systems (e.g., [24]).

<sup>3</sup>In the terminology of multiple hypothesis tests, the false-alarm probability corresponds to the family-wise error rate [26].

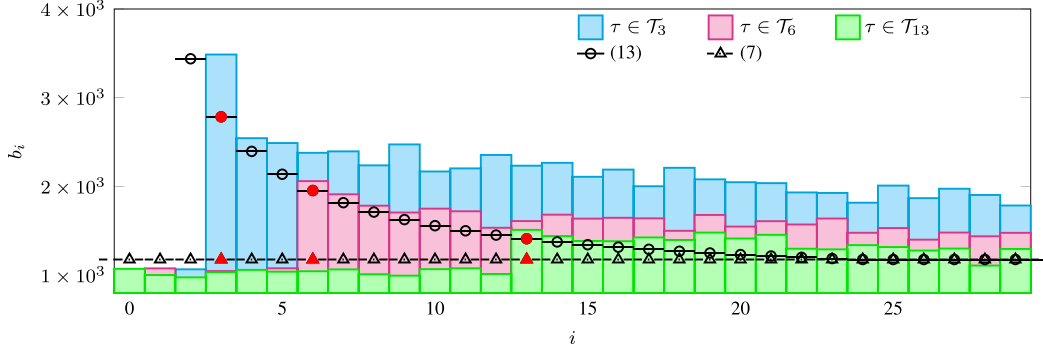


Fig. 2. Example of energy bin vectors and threshold profiling. Each element  $b_i$  is an instantiation of the RV  $\mathbf{b}_i$ . The true bins containing the first signal samples are highlighted with red marks.

### A. Constant Threshold Profile

The most common approach in solving the hypothesis testing problem (6) is to set a constant threshold profile (i.e.,  $\xi_i = \xi_{\text{fa}} \forall i \in \mathcal{B}$ , also known as CFAR) such that  $P_{\text{fa}}(\boldsymbol{\theta}) = P_{\text{fa}}^*$ , where  $P_{\text{fa}}^*$  is the maximum tolerable false-alarm probability. Since the false-alarm event occurs in the absence of the signal, the energy bins  $\mathbf{b}_i$  contain only noise and are independent and identically distributed RVs. From (6a), the threshold  $\forall i \in \mathcal{B}$  is

$$\xi_{\text{fa}} = F_{\mathbf{b}_i}^{-1} \left( (1 - P_{\text{fa}}^*)^{\frac{1}{N_{\text{bin}}}} \mid \boldsymbol{\theta}_{\mathbf{b}} \right). \quad (7)$$

### B. Variable Threshold Profile

We propose the use of a variable threshold profile that fulfills a requirement on the ranging error. In particular, the threshold profile is set to guarantee a minimum tolerable value  $P_c^*$  for the probability of selecting the correct bin index (i.e.,  $i$  when  $\tau \in \mathcal{T}_i$ ) to minimize the ranging error.

*Lemma 3.1:* The probability of selecting the correct bin for  $\tau \in \mathcal{T}_i$  is lower bounded by

$$P_{c,i}(\boldsymbol{\theta}_{\mathbf{b}}) \geq P_c^* \quad (8)$$

when  $\xi_j^{(L)} \leq \xi_j \leq \xi_j^{(U)}, \forall j \in \mathcal{B}$ , where

$$\xi_j^{(L)} = \min_{\boldsymbol{\theta}: \lambda_j = 0} F_{\mathbf{b}_j}^{-1} \left( P_c^*^{\frac{1}{N_{\text{bin}}}} \mid \boldsymbol{\theta}_{\mathbf{b}} \right) \quad (9a)$$

$$\xi_j^{(U)} = \max_{\boldsymbol{\theta}: \tau \in \mathcal{T}_j} F_{\mathbf{b}_j}^{-1} \left( 1 - P_c^*^{\frac{1}{j+1}} \mid \boldsymbol{\theta}_{\mathbf{b}} \right). \quad (9b)$$

*Proof:* For  $\tau \in \mathcal{T}_i$ , it is

$$\begin{aligned} \mathbf{b}_i \frac{N_{\text{sr}}}{\sigma^2} \mid \boldsymbol{\theta}_{\mathbf{b}} &\sim \chi_{N_{\text{sr}} N_{\text{sb}}}^2(\lambda_i) \text{ with } \lambda_i \neq 0 \\ \mathbf{b}_j \frac{N_{\text{sr}}}{\sigma^2} \mid \boldsymbol{\theta}_{\mathbf{b}} &\sim \chi_{N_{\text{sr}} N_{\text{sb}}}^2(0) \forall j \in \mathcal{I}_i(i). \end{aligned} \quad (10)$$

Since  $F_{\mathbf{b}_j}(\xi_j)$  is monotonically increasing with  $\xi_j$ , if  $\xi_j^{(L)} \leq \xi_j \leq \xi_j^{(U)} \forall j \in \mathcal{B}$ , then

$$P_{c,i}(\boldsymbol{\theta}_{\mathbf{b}}) \geq \left[ 1 - F_{\mathbf{b}_i}(\xi_i^{(U)} \mid \boldsymbol{\theta}_{\mathbf{b}}) \right] \prod_{j \in \mathcal{I}_i(i)} F_{\mathbf{b}_j}(\xi_j^{(L)} \mid \boldsymbol{\theta}_{\mathbf{b}}) \quad (11a)$$

$$= P_c^*^{\frac{1}{i+1}} P_c^*^{\frac{i}{N_{\text{bin}}}} \quad (11b)$$

where the first term in (11b) is from (9b) and the second term is from (9a). As  $P_c^* \in [0, 1]$ , then for all  $i \in \mathcal{B}$ , we have

$$P_{c,i}(\boldsymbol{\theta}_{\mathbf{b}}) \geq P_c^*^{\frac{1}{i+1}} P_c^*^{\frac{i}{i+1}} \quad (12)$$

resulting in (8).  $\square$

Lemma 3.1 enables the design of a variable threshold profile for fulfilling a requirement on ranging error. An approach that accounts for requirements on both false-alarm probability and ranging error sets the variable threshold profile as

$$\xi_i = \max \{ \xi_{\text{fa}}, \xi_i^{(U)} \}. \quad (13)$$

Then, (13) fulfills both the false-alarm probability requirement and the requirement on ranging error when  $\xi_{\text{fa}} \leq \xi_i^{(U)}, \forall i \in \mathcal{B}$ .

Fig. 2 shows some examples of energy bin vectors for three different TOAs. The energy values are compared with the constant (7) and variable (13) threshold profiles with  $P_{\text{fa}}^* = 0.01$  and  $P_c^* = 0.9$ . It can be observed that the variable threshold profile is expected to provide a lower false-alarm probability since  $\xi_i \geq \xi_{\text{fa}} \forall i \in \mathcal{B}$ . Moreover, incorrectly selecting a bin is less probable, which is beneficial for reducing early-detection probability (the probability of estimating a smaller TOA than the true one) and ranging error.

Note that (9a) and (9b) require to invert the CDF of the  $j$ th bin, which is a noncentral chi-square RV. This is a complex operation often performed through numerical evaluations or via approximations. For this reason, a tractable model is necessary for the application of Lemma 3.1 in practical scenarios.

### C. Tractable Model for Variable Threshold Profile

The design of variable threshold profile described in Section III-B needs to consider the following two aspects: i) the noncentral chi-squared RV does not admit a closed-form CDF expression; and ii) the instantaneous channel state information  $\boldsymbol{\theta}$  is usually not perfectly known. In particular, a tractable model is proposed for the variable threshold profile using an approximation for the CDF of the bin value conditioned on the channel statistics  $\bar{\boldsymbol{\theta}}$  instead of the channel instantiation  $\boldsymbol{\theta}$ . Specifically, we consider the noncentrality parameter as  $\lambda_i \simeq \bar{\lambda}_i$ , which is deterministic and is calculated from (2) through the average channel parameters  $\bar{\boldsymbol{\theta}}$ . When  $\lambda_i$  is replaced with  $\bar{\lambda}_i$  in (1), the

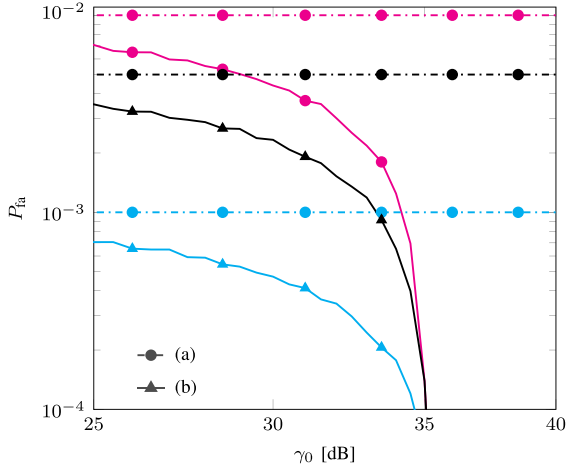


Fig. 3. False-alarm probability as a function of  $\gamma_0$  for  $P_{fa}^* = 10^{-2}$  (magenta),  $P_{fa}^* = 5 \times 10^{-3}$  (cyan),  $P_{fa}^* = 10^{-3}$  (black) with (a)  $\xi_i = \xi_{fa}$  and (b)  $\xi_i = \max\{\xi_{fa}, \xi_i^{(U)}\}$ .

chi-squared RVs can be approximated by Gaussian RVs [25]. In particular, the  $N_{bin}$  hypothesis tests in (6) for a given  $\hat{\theta}$  become

$$\begin{aligned} \tilde{\mathcal{H}}_{0i} &: b_i \frac{N_{sr}}{\sigma^2} \bar{\theta}_b \mathcal{N}(N_{sr}N_{sb}, 2N_{sr}N_{sb}) \\ \tilde{\mathcal{H}}_{1i} &: b_i \frac{N_{sr}}{\sigma^2} \bar{\theta}_b \mathcal{N}(N_{sr}N_{sb} + \bar{\lambda}_i, 2N_{sr}N_{sb} + 4\bar{\lambda}_i). \end{aligned}$$

From (9), a closed-form expression for  $\xi_i^{(L)}$  and  $\xi_i^{(U)}$  is

$$\begin{aligned} \xi_i^{(L)} &\simeq \Phi^{-1}(P_c^{\frac{1}{N_{bin}}}) \sqrt{2N_{sr}N_{sb} + N_{sr}N_{sb}} \\ \xi_i^{(U)} &\simeq \Phi^{-1}(1 - P_c^{\frac{1}{i+1}}) \sqrt{2N_{sr}N_{sb} + 4\bar{\lambda}_i + N_{sr}N_{sb} + \bar{\lambda}_i} \end{aligned}$$

where  $\Phi^{-1}(\cdot)$  is the inverse CDF of a standard Gaussian RV.

A further simplification for  $\xi_i^{(U)}$  can be made for tapped-delay line channels with exponential power dispersion profile [27]. In such a case, the average noncentrality parameter for the  $i$ th bin  $\bar{\lambda}_i$  when  $\tau \in \mathcal{T}_i$  can be expressed as

$$\bar{\lambda}_i = \bar{\nu}_0 \left( \frac{d_i}{\delta_0} \right)^{-\alpha}$$

where  $\bar{\nu}_0$  is the noncentrality parameter corresponding to the correct bin when the first path is related to the reference distance  $d = \delta_0$ ,  $\alpha$  is the path-loss exponent,  $d_i = g(i)c$ , and  $c$  is the speed of light. The tractable expression for  $\xi_i^{(U)}$  is found to be

$$\begin{aligned} \xi_i^{(U)} &= \Phi^{-1}(1 - P_c^{\frac{1}{i+1}}) \sqrt{2N_{sr}N_{sb} + 4\bar{\nu}_0 \left( \frac{d_i}{\delta_0} \right)^{-\alpha}} \\ &+ N_{sr}N_{sb} + \bar{\nu}_0 \left( \frac{d_i}{\delta_0} \right)^{-\alpha}. \end{aligned} \quad (15)$$

#### IV. CASE STUDY

The performance improvement of the variable threshold profile is now evaluated using TCS algorithms compared to the constant threshold profile. The ED is set with  $N_{sr} = 128$ ,  $T_{obs} = T_{sr} = 100$  ns,  $T_d = 2$  ns, and  $N_{bin} = 50$ . A tapped-delay line channel is considered, with exponential power dispersion

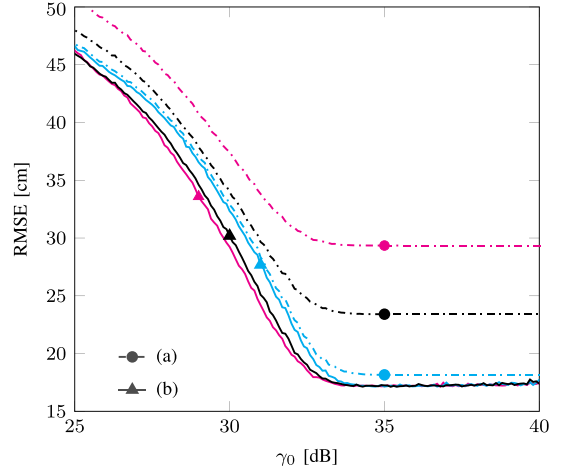


Fig. 4. Root-mean-square of the ranging error as a function of  $\gamma_0$  for  $P_{fa}^* = 10^{-2}$  (magenta),  $P_{fa}^* = 5 \times 10^{-3}$  (cyan),  $P_{fa}^* = 10^{-3}$  (black) with (a)  $\xi_i = \xi_{fa}$  and (b)  $\xi_i = \max\{\xi_{fa}, \xi_i^{(U)}\}$ .

profile and  $\alpha = 2.4$  as in [27]. Results are presented for different values of  $\gamma_0$ , which is the SNR at the reference distance  $\delta_0 = 1$  m. Two different methods for choosing thresholds  $\xi_i, \forall i \in \mathcal{B}$ , are considered: the constant threshold profile  $\xi_i = \xi_{fa}$  and the variable threshold profile  $\xi_i = \max\{\xi_{fa}, \xi_i^{(U)}\}$  through (15) with  $P_c^* = 0.9$  (i.e., correct bin selected at least 90% of times).

Fig. 3 shows the false-alarm probability as a function of  $\gamma_0$  for  $P_{fa}^* = 10^{-3}, 5 \times 10^{-3}$ , and  $10^{-2}$ . It can be observed that the false-alarm probability obtained with the variable threshold profile is significantly lower than or equal to that obtained with the constant threshold profile, and it decreases with  $\gamma_0$ .<sup>4</sup>

Fig. 4 shows the root-mean-square of the ranging error  $|\hat{\tau} - \tau|$  as a function of  $\gamma_0$  for  $P_{fa}^* = 10^{-3}, 5 \times 10^{-3}$ , and  $10^{-2}$ , respectively. It can be observed that the variable threshold profile always outperforms the constant threshold profile and the performance gap increases with  $\gamma_0$  and  $P_{fa}^*$ . The floor value obtained with the variable threshold profile corresponds to the minimum root-mean-square error that is achievable through an ED in ideal propagation conditions. In particular, such a floor is  $cT_d/\sqrt{12} \simeq 17$  cm, which is the quantization error due to the temporal resolution of the bin vector.

#### V. CONCLUSION

A variable threshold profile for TOA-based ranging was proposed, where the ED threshold design is driven by requirements on both ranging error and false-alarm probability. Tractable expressions for the variable threshold profile were obtained, enabling a simple and yet efficient ranging system design. In particular, the design of the variable threshold profile is based only on statistics rather than instantiations of the wireless channel. A case study for a wideband ranging system showed that the variable threshold profile can significantly improve the performance, compared to conventional methods, in terms of both ranging error and false-alarm probability.

<sup>4</sup>Although not reported in detail for brevity, it was observed that the variable threshold profile improves the false-alarm probability with respect to the constant threshold profile without any degradation of the detection probability.

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